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Quantification of microscopic diffusion anisotropy disentangles effects of orientation dispersion from microstructure: Applications in healthy volunteers and in brain tumors

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Abstract

The anisotropy of water diffusion in brain tissue is affected by both disease and development. This change can be detected using diffusion MRI and is often quantified by the fractional anisotropy (FA) derived from diffusion tensor imaging (DTI). Although FA is sensitive to anisotropic cell structures, such as axons, it is also sensitive to their orientation dispersion. This is a major limitation to the use of FA as a biomarker for “tissue integrity”, especially in regions of complex microarchitecture. In this work, we seek to circumvent this limitation by disentangling the effects of microscopic diffusion anisotropy from the orientation dispersion. The microscopic fractional anisotropy (μFA) and the order parameter (OP) were calculated from the contrast between signal prepared with directional and isotropic diffusion encoding, where the latter was achieved by magic angle spinning of the q-vector (qMAS). These parameters were quantified in healthy volunteers and in two patients; one patient with meningioma and one with glioblastoma. Finally, we used simulations to elucidate the relation between FA and μFA in various micro-architectures.

Generally, μFA was high in the white matter and low in the gray matter. In the white matter, the largest differences between μFA and FA were found in crossing white matter and in interfaces between large white matter tracts, where μFA was high while FA was low. Both tumor types exhibited a low FA, in contrast to the μFA which was high in the meningioma and low in the glioblastoma, indicating that the meningioma contained disordered anisotropic structures, while the glioblastoma did not. This interpretation was confirmed by histological examination.

We conclude that FA from DTI reflects both the amount of diffusion anisotropy and orientation dispersion. We suggest that the μFA and OP may complement FA by independently quantifying the microscopic anisotropy and the level of orientation coherence.

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Introduction

The most established technique for non-invasive investigations of the microstructure of the central nervous system is diffusion tensor imaging (DTI) (Basser et al., 1994). DTI provides a means of estimating the rate of diffusion and the diffusional anisotropy in tissue, frequently expressed in terms of the mean diffusivity (MD) and the fractional anisotropy (FA), respectively. The diffusion anisotropy has been shown to correlate with the progression of a wide variety of conditions (Kubicki et al., 2002). For example, reduced FA is observed during aging (Hsu et al., 2010; Sullivan and Pfefferbaum, 2006), and in neurodegenerative diseases such as dementia (Englund et al., 2004; Santillo et al., 2013), Parkinson’s disease (Surova et al., 2013), Alzheimer’s disease (Sjobeck et al., 2010), and multiple sclerosis (Rovaris et al., 2005). By contrast, the value of FA tends to increase during white matter (WM) maturation (Lebel...
et al., 2008; Löbel et al., 2009) and after specific forms of training, such as juggling (Scholz et al., 2009).

While FA is closely sensitive to microstructural alterations, such as demyelination, it also reflects a wide variety of non-specific and possibly confounding effects. One of the most prominent confounders of FA is the partial volume effect (PVE). Partial volume effects are especially relevant for diffusion-MRI (dMRI) where voxel volumes are typically on the scale of ~10 mm³, resulting in a high probability for the MR signal to originate from water residing in different types of tissue. This includes voxels that are located at the interface between nerve bundles with different orientation, and at the interface between brain tissue and cerebrospinal fluid (CSF). Thus, the signal from individual voxels frequently reflects an average of different diffusion profiles. This invariably leads to less pronounced diffusion directionality, i.e., lower FA (Oouchi et al., 2007; Westin et al., 2002). Consequently, FA correlates with structure size since smaller structures include a larger fraction of voxels that interface with surrounding tissue than larger structures (Szczepankiewicz et al., 2013; Vos et al., 2011). Another aspect of PVE is the presence of crossing, kissing, fanning, and other irregular WM geometries within a voxel, which reduce the FA by inducing a higher degree of orientation dispersion (Alexander et al., 2001; Nilsson et al., 2012). Thus, the utility of FA as a biomarker in regions of complex WM architecture is impeded because it entangles multiple effects into a single value. Although frequently overlooked, this is not an idle theoretical issue but has practical consequences. For example, elevated values of FA have been found in crossing fibers in patients with Alzheimer’s disease (Douaud et al., 2011; Teipel et al., 2014). This seemingly counter-intuitive result is explained by the selective damage to one of the fiber populations in the region (Douaud et al., 2011), resulting in reduced orientation dispersion and thus elevated FA. It is also worth noting that FA is an intrinsically poor biomarker in gray matter (GM) due to the high orientation dispersion of neurites in the cortex (Shemesh and Cohen, 2011). Thus, reliable use of FA may be confined to regions of highly coherent WM (De Santis et al., 2013), which is estimated to account for less than 10% of the total white matter of the human brain (Vos et al., 2012). This has prompted the search for methods that accurately model microscopic changes in complex neural tissue.

It has been shown that the effects of orientation and restriction can be disentangled by extending the conventional single pulsed-field-gradient (sPFG) experiment (Stejskal and Tanner, 1965) to include double, or multiple, pulsed-field-gradients (dPFG and mPFG, respectively) (Mitra, 1995). In dPFG experiments information can be derived from the dependence of the signal amplitude on the angle between two successive encoding blocks. Several methods have been proposed for the quantification of microscopic anisotropy from such data. To this end, Lawrenz and Finsterbusch (2013) used a fourth-order tensor parameterization suggested by Lawrenz et al. (2010) to map the microscopic diffusion anisotropy in human white matter in vivo. Jespersen et al. (2013) developed a rotationally invariant dPFG encoding scheme and mapped the microscopic anisotropy in an excised monkey brain in terms of the fractional eccentricity.

Recently, Lasić et al. (2014) formulated a framework for the quantification of microscopic diffusion anisotropy and orientation dispersion in terms of the microscopic fractional anisotropy (μFA) and order parameter (OP), respectively. These parameters were derived from the contrast between the signal acquired in diffusion weighting (DW) experiments that used conventional diffusion encoding as well as isotropic encoding based on magic angle spinning of the q-vector (qMAS) (Eriksson et al., 2013). Briefly, magic angle spinning is an established NMR spectroscopy method where a sample is rotated around its own axis at a specific angle relative to the B0-field to minimize the influence of chemical shift anisotropy on the observed NMR spectrum. In qMAS, harmonic gradient modulation is used to create a q-vector that performs a precession at the magic angle in order to exert equal diffusion encoding in all spatial directions while the sample remains stationary.

Although isotropic encoding can be achieved by combining multiple trapezoidal encoding blocks (Butts et al., 1997; Wong et al., 1995), the qMAS technique offers a time efficient gradient modulation scheme (Toppgaard, 2013). The qMAS-encoded signal attenuation becomes independent of contributions from anisotropic diffusion, and is sensitive only to the rate of isotropic diffusion (Eriksson et al., 2013). As a proof-of-principle, Lasić et al. (2014) implemented the qMAS technique on a NMR spectrometer and a clinical scanner, showing that microscopic anisotropy could be detected in phantoms that contained ordered and disordered anisotropic micro-domains.

In this work we performed the first in vivo experiments using qMAS diffusion encoding, and we parameterize the microscopic anisotropy of the human brain based on the framework presented by Lasić et al. (2014). We also demonstrated the feasibility of quantifying microscopic anisotropy in a clinical setting by using it to infer information on tissue structure in two types of brain tumors. Finally, we compared the results to simulated data to elucidate how the measures of anisotropy respond to various changes in micro-architecture, and expanded on the possibilities to use this novel method in clinical research to access information that is unavailable when using conventional methods.

Theory

In conventional DTI, the diffusion on the voxel scale is assumed to be Gaussian and is described by a rank-2 tensor (D) (Basser et al., 1994). The same description can be employed at a sub-voxel scale; meaning that each coherent segment of the underlying microgeometry can be considered as a domain in which the diffusion is Gaussian and described by a domain diffusion tensor (Dk). The voxel scale tensor can be described as the average of all domain tensors, according to Eq. (1)

\[
D = \langle D_k \rangle,
\]

where \(D = D_k\), only when the voxel contains identical domains that are perfectly aligned. In all other cases \(D\) will depend on the distribution of domain tensor eigenvalues, and their orientation (Fig. 1). Here, we denote objects pertaining to microscopic domains by a subscript ‘k’. Consider three common parameterizations of \(D\): the mean diffusivity (MD), the variance of the diffusion tensor eigenvalues (\(V_k\)), and the fractional anisotropy (FA), defined in Eqs. (2), (3), and (4) respectively (Basser and Pierpaoli, 1996)

\[
MD = \mu(D) = \frac{\text{Tr}(D)}{3},
\]

\[
V_k = \sigma(D) = \text{Var}(\text{Eig}(D)),
\]

\[
FA = \mathcal{A}(\text{MD}, V_k) = \frac{3}{2} \left(1 + \frac{\text{MD}^2}{V_k} \right)^{-\frac{3}{2}}.
\]

Note that \(\mu(D_k), \sigma(D_k)\) and \(\mathcal{A}(\text{MD}, V_k)\) yield the corresponding parameters for a single domain, denoted \(\text{MD}_k, V_{k} \) and \(\text{FA}_k\), respectively. From Eq. (1) to Eq. (4), it is clear that the FA represents the amount of microscopic anisotropy that persists to the voxel scale and is determined by the coherence of the domain orientations (Westin et al., 2002). To circumvent this dependency, Lasić et al. (2014) suggested a method to measure the microscopic anisotropy in terms of the microscopic fractional anisotropy (μFA). Conceptually, in a system of identical and parallel domains the diffusion
anisotropy of each domain will persist to the voxel scale, rendering $\mu FA = FA = FA_k$ (Fig. 1A). By contrast, randomly oriented domains exhibit isotropic voxel scale diffusion, rendering $FA = 0$, however, the microscopic anisotropy is unaffected by the orientation dispersion and thus $\mu FA = FA_k$ (Fig. 1C).

It should be clear that individual domains cannot be probed directly using conventional DTI. Instead, the microscopic anisotropy can be inferred from the amount by which the diffusion weighted signal deviates from monoexponential attenuation, commonly referred to as the diffusional kurtosis (Jensen et al., 2005). However, kurtosis is not specific to microscopic anisotropy since it is also sensitive to the presence of multiple diffusion coefficients. Further, Mitra (1995) showed that these two effects cannot be distinguished in a conventional sPFG experiment, but that it could be done using dPFG experiments. Here, we separate the two effects by using the contrast between conventional and isotropic diffusion encoding (Lasić et al., 2014). The concept is understood by considering the MR signal ($S$) as a function of the magnitude of diffusion encoding ($b$), and the distribution of diffusion coefficients ($P$), according to Eq. (5)

$$S_n(b) = S_0 \int P(D|N) \cdot e^{-|bD|} dD,$$

where $P(D|N)$ reads as the probability distribution of diffusion coefficients when employing the encoding tensor $N$, and $D = N \cdot D$, where ‘:’ denotes the double inner product. The encoding tensor is introduced to facilitate the analysis of both conventional and isotropic encoding (Westin et al., 2014). Conventional diffusion encoding is anisotropic, i.e., the diffusion sensitizing gradient is employed in one specific direction $n$, where $n = [n_x, n_y, n_z]$ and $|n| = 1$. The corresponding encoding tensor is defined as $N = nn^T$ ($3 \times 3$ matrix with a single non-zero eigenvalue), and the b-matrix is given by $B = b \cdot N$ (Basser et al., 1994).

For low to moderately high $b$-values, the signal described in Eq. (5) mainly depends on the expected value and the variance of the distribution of diffusion coefficients. The expected value, or first moment, of $P$ is reflected in the initial slope of the signal attenuation, and is equal to the apparent diffusion coefficient in the direction defined by $N$, according to $ADC = E[P(D|N)]$. The variance, or second central moment, of $P$ is reflected in the departure of the signal attenuation from monoexponentiality, and is related to the apparent diffusion kurtosis ($K$) mapped in DKI, such that $Var(P(D|N)) = K \cdot ADC^2/3$ (Jensen et al., 2005).

The dependence of the distribution of diffusion coefficients on $N$ is essential to understanding the calculation of the microscopic anisotropy. We highlight this dependence by considering an ideal system that contains an ensemble of anisotropic domains that are randomly oriented and axially symmetric, i.e., the system is rotationally invariant and all domain tensors are defined by two eigenvalues. This system is anisotropic on the microscopic scale, but isotropic on the voxel scale, hence $FA = 0$. However, the microscopic anisotropy can be recovered from the variance of the distribution of diffusion coefficients reflected in the departure from monoexponential signal attenuation. In the ideal system, the average variance of the domain tensor eigenvalues ($\langle V_{\lambda,k} \rangle$) is related to the variance of the distribution of diffusion coefficients ($V_a$) according to Eq. (6) (Lasić et al., 2014)

$$\langle V_{\lambda,k} \rangle = \frac{5}{2} V_a,$$

where $V_a = Var(P(D|N))$. The subscript ‘a’ indicates that the variance is induced only by the presence of anisotropy. The microscopic fractional anisotropy is defined by substituting $V_a$ in Eq. (4) with the right hand side of Eq. (6), according to Eq. (7) (Lasić et al., 2014; Topgaard and Lasić, 2013)

$$\mu FA = \mathcal{A} \left( \frac{5}{2} V_a \right) = \frac{3}{2} \left( 1 + \frac{MD^2}{\frac{5}{2} V_a} \right)^{-\frac{1}{2}}.$$

The definition in Eq. (7) was originally suggested by Topgaard and Lasić (2013), but an analogous parameter, the fractional eccentricity (FE), was independently developed by Jespersen et al. (2013). Note that the $\mu FA$ and FE differ only by a constant factor such that $\mu FA = \sqrt{3/2} \cdot FE$ (Jespersen et al., 2014a,b; Lasić et al., 2014).
Applying Eq. (7) to an ideal system is able to perfectly describe the μFA as an analog to FA that is not sensitive to the effects of orientation dispersion (Fig. 1). However, assumptions made in the ideal system may not be valid in biological tissue. In such cases, the μFA can still be quantified by relaxing the demands of the ideal system and compensating for the introduced error. Here we consider departure from rotation invariance, and the presence of multiple sources of variance.

Rotation invariance can be achieved by constructing the powder average of the signal and is required in systems that exhibit residual anisotropy (FA > 0). The powder average is the arithmetic average of the signal across multiple rotations of the diffusion encoding gradients, and will render a signal that is insensitive to rotations of the object. Here we denote the powder averaged signal and distribution function as $S$ and $P$, respectively. Note that the expected value of the powder averaged distribution yields the mean diffusivity, i.e., $\langle P(D)N \rangle = \mu_D$. 

Variance in the distribution of diffusion coefficients can be a consequence of both anisotropy and presence of multiple isotropic components. This is relevant for the evaluation of Eq. (7) where only the variance arising due to the presence of microscopic anisotropy is considered. Thus, in cases where all domains cannot be assumed to have equal isotropic diffusivity, i.e., the domains have different MDs, the contribution to total variance ($V_t$) from isotropic components ($V_i$) must be quantified and removed, according to Eq. (8)

$$V_a = V_t - V_i.$$

To calculate $V_a$ according to Eq. (8) we must find an independent means of measuring $V_i$ and $V_t$. We know from DKI that $V_i$ can be quantified by performing a conventional diffusion experiment, according to $V_i = \text{Var}(PD(N))$. Since $P$ is affected not only by the underlying microenvironment, but also by the shape of the encoding tensor, $V_i$ can be quantified by employing isotropic diffusion encoding that is designed to exert equal encoding strength in all spatial directions in a single preparation of the signal. We define the isotropic encoding tensor (1, 3 × 3 matrix) as one-third of the identity matrix so that all its eigenvalues are equal, and $\text{Tr}(I) = 1$. This mode of encoding is insensitive to the domain orientations, and if the diffusion is approximately Gaussian, it is rotationally invariant and independent of microscopic anisotropy. Note that when isotropic encoding is used, $P$ and $P$ are interchangeable since $I$ has no defined direction. For isotropic encoding the signal in Eq. (5) is a function of $P(D)$ which denotes the distribution of domain mean diffusivities since $I \cdot D = MD$. The remaining variance is due to heterogeneous domain mean diffusivities, and is defined as $V_t = \text{Var}(PD(N))$. In summary, anisotropic and isotropic diffusion encoding at sufficiently high $b$-values can be used to quantify $V_i$ and $V_t$, respectively. The μFA can then be calculated according to Eqs. (7) and (8).

Finally, we note that the interpretation of $V_i$ in Eq. (8) is valid if the two probability distribution functions are related in terms of a convolution, according to $P(D|N) = R(D) \ast P(D|I)$ (see Fig. 2), where $R(D)$ is the anisotropy response function and $V_a = \text{Var}(R(D))$, according to probability theory and the arithmetic of random variables. Thus, the analysis assumes that the variance of the anisotropy response function is equal for all domains. This assumption may be invalid, for example, in mixtures of WM and CSF where the anisotropy response functions are expected to be markedly different. The effects of such unfavorable conditions on the validity of μFA calculations are investigated in the Simulation experiments.

**Methods**

**Imaging protocols**

Data was acquired using a Philips Achieva 3T system, equipped with 80 mT/m gradients with a maximum slew rate of 100 mT/m/ms on axis, and an eight channel head coil.

The in vivo experiment was designed to evaluate the validity of the μFA model and was therefore acquired using a high $b$-value sampling rate, employing ten equidistant $b$-values between 100 and 2800 s/mm$^2$. Thereby, the sequence was limited to five image slices. Each set of data (one set per subject) contained images prepared with both the isotropic qMAS and harmonically modulated anisotropic encoding (Fig. 3). Harmonic modulation is preferred to trapezoidal encoding to ensure equal diffusion times for both types of encoding (Eriksson et al., 2013). All DW data were acquired using an echo time of 160 ms, repetition time of 2000 ms, 96 × 96 acquisition matrix, spatial resolution of 3 × 3 × 3 mm$^3$, partial Fourier factor of 0.8, and a SENSE factor of 2. Regardless of encoding technique, each encoding block, before and after the 180°-pulse, lasted 62.5 ms. Anisotropic encoding was performed in 15 directions for each $b$-value using harmonically modulated gradients according to Lasić et al. (2014). The directions were distributed using an electrostatic repulsion scheme (Jones et al., 1999). The isotropic encoding was repeated 15 times per $b$-value. This resulted in equal amounts of images and scan time for both techniques. The combined scan time for the isotropic and anisotropic encoding sequences was 10:12 min.

![Fig. 2. Schematic example of the distribution of diffusion coefficients when employing encoding that is isotropic (left, $P(D|I)$) and anisotropic (right, $P(D|N)$). The convolution visualizes how the variance of $P(D|I)$ is added to the variance of the anisotropy response function $R(D)$, rendering the total variance in $P(D|N)$. This example depicts a system that contains axially symmetric and randomly oriented domains where $MD = 0.75 \pm 0.05 \mu$m$^2$/ms, and the axial and radial domain diffusion is $AD = MD + 1.0 \mu$m$^2$/ms and $RD = MD - 0.5 \mu$m$^2$/ms, respectively (middle panel). Thus, the variance of the anisotropy response function is equal for all domains. The fact that the system contains anisotropic domains is reflected in the width of $R(D)$, indicating that there is a difference between the eigenvalues of the domain tensors. The prolate symmetry of the domain tensors can be discerned from the shape of $R(D)$, where the slow diffusion ($RD$) is the most probable while the fast diffusion ($AD$) is the least probable (Eriksson et al., 2013). Note that the area under each distribution equals unity, and that the y-axes have been adjusted to improve legibility.](image-url)
data employing anisotropic encoding for encoding strengths in Eq. (9) corresponds to and curvature of the signal attenuation function, respectively. Note that the diffusion time and the time required for the diffusing medium to reach the parameterization should not affect the mean diffusivity unless the diffusion medium is constrained to the range between the total variance and zero.

Post-processing and parameterization

Motion correction and eddy-current correction was applied to DWI data using Elastix (Klein et al., 2010). The first moment, and the second central moment of the distribution of diffusion coefficients was estimated by regressing the inverse Laplace transform of the gamma distribution function onto the acquired signal (Lasie et al., 2014; Roding et al., 2012). The signal was modeled, according to Eq. (9)

\[ S(b) = S_0 \left( 1 + \frac{b \cdot V}{MD} \right)^{\frac{MT}{MD}} \]  

where MD and V were the fitting variables representing the initial slope and curvature of the signal attenuation function, respectively. Note that V in Eq. (9) corresponds to \( V_i \) and \( V_l \) when the model is regressed onto data from the powder averaged anisotropic and isotropic diffusion encoding experiments, respectively. Three constraints were introduced in the fitting procedure to eliminate non-physical results. First, the MD was constrained to be equal in the two acquisitions by assuming that \( \langle P(D|N) \rangle = \langle P(D|I) \rangle = MD \). This assumption is reasonable since the choice of encoding technique should not affect the mean diffusivity unless the diffusion time and the time required for the diffusing medium to reach the relevant restrictions are at the same scale, which is rarely the case for DWI in vivo (Nilsson et al., 2009, 2013). Second, \( V_i \) was constrained to the range between the total variance and zero (\( V_i \geq V_l \geq 0 \)). Finally, signal that was attenuated below 5% (\( S(b) < 0.05 \cdot S_0 \)) was excluded from the fitting procedure. This was done to avoid detection of false variance in regions where a strong diffusion weighting rendered a signal that was elevated due to the noise floor. This is expected to affect only voxels where MD > 1.1 \( \mu m^2/\text{ms} \).

FA was calculated through conventional DTI analysis from the data employing anisotropic encoding for encoding strengths \( b \leq 1000 s/mm^2 \). The \( \mu FA \) was calculated according to Eq. (8). Finally, the orientation coherence of the domains was quantified by the order parameter which is a well-established parameter for describing the order in liquid crystals. It is defined as \( OP = \langle 3 \cos^2(\theta_b) - 1/2 \rangle \), where \( \theta_b \) is the angle between the domain and voxel scale symmetry axes. Thus, the OP provides a measure of orientation dispersion that has a simple geometric interpretation where \( OP = 1 \) indicates perfectly coherent alignment and \( OP = 0 \) indicates randomly oriented domain orientations. The OP can also be calculated from the microscopic and voxel scale variance, according to Eq. (10) (Lasie et al., 2014)

\[ OP = \frac{V_A}{\sqrt{\sum_k V_{\lambda_k}}} \]  

Note that OP is not equivalent to the orientation dispersion index used in NODDI (Zhang et al., 2012), and that it can be calculated for any given orientation distribution function.

In vivo experiments

Imaging was performed on eight healthy volunteers (age 32 ± 4 years, all male) and two patients with brain tumors (one female, 62 years, with meningioma, WHO grade I; and one male, 46 years, with glioblastoma, WHO grade IV). Written consent was obtained from all subjects and the study was approved by the Regional Ethical Review Board at Lund University.

Analysis of diffusion parameters was performed at the group level, as well as in a single representative subject. Three regions of interest (ROI) were selected in the WM; the splenium of the corpus callosum (CC), the corticospinal tract (CST), and the anterior crossing region (CR) where frontal projection fibers from the genu of the corpus callosum and thalamic radiation of the internal capsule intersect (see Assaf and Pasternak 2008). One ROI was also placed in the superior part of the thalamus (THA), which contains a mixture of WM and GM. The ROIs were delineated manually, using MD, FA and \( \mu FA \) maps for guidance; the operator was instructed to avoid voxels that contained GM or CSF.

The healthy individual was investigated with respect to the signal parameterization and parameter distribution in all four ROIs. One additional ROI was placed in the lateral ventricles to investigate the signal attenuation in the isotropic and rapidly diffusing CSF. The analysis of the parameter distribution was based on the ROIs while the signal and model fit was inspected in a single voxel in each ROI. Further, the voxel-wise correlation between combinations of FA, \( \mu FA \) and OP were...
evaluated. This analysis was performed in one axial slice of the image volume and the parameter maps were masked to remove interference from irrelevant regions of the head. The strength of the association was quantified by the coefficient of determination ($r^2$, Pearson's linear correlation coefficient squared).

The healthy volunteer group was investigated with respect to the parameter distribution in the CC, CST, CR and THA. In order to elucidate if the three WM regions were different with respect to parameter mean values, F-tests (one-way ANOVA, assuming independent samples) were performed on the distributions of MD, FA, µFA, OP, $V_1$ and $V_2$ in the CC, CST and CR. The threshold for significance was set at $\alpha = 0.05/6$ (Bonferroni correction for six tests).

The tumors were compared with respect to their FA and µFA by placing ROIs in one axial slice through each tumor. The ROIs were defined manually and the inclusion of WM, GM and CSF was avoided. Both tumors were resected one day after the MRI procedure and histological evaluation of the tumors was performed, in accordance with local clinical routine. Each tumor specimen was fixed in 4% buffered formaldehyde solution, embedded in paraffin, and sectioned at 4 µm. The sections were stained with hematoxylin–eosin in order to visualize the tissue structure and cell morphology. Microscopy was performed on an Olympus BX50. The cell shape and presence of tissue fascicles was investigated qualitatively and compared to corresponding diffusion parameters. Finally, structure tensor analysis (Peyré, 2011) was performed on the microphotos to enhance the visibility of cell structure orientations.

Simulation experiments

Simulation experiments were performed to investigate the qualitative behavior of FA and µFA in scenarios where the underlying system contained complex diffusion profiles. These scenarios were designed to mimic a range of effects that may be found in experimental data. The results were evaluated in terms of the value, effect size, effect direction, and accuracy of the FA and µFA.

The simulations included three types of model components (C) with varying water fractions ($f_i$). The first component was designed to represent the anisotropic diffusion in WM ($C_0$). For simplicity, all anisotropic domains were assumed to be axially symmetric and were described by their radial (RDk) and axial diffusivity (ADk). These were set to ADk = 1.7 and RDk = 0.2 µm²/ms. The orientation dispersion was modeled with the Watson distribution (Sra and Karp, 2013; Zhang et al., 2011) where the concentration parameter ($\kappa$) is related to the order parameter according to Eq. (11)

$$\text{OP} = \frac{1}{2} \mathcal{M}(3/2, 5/2, \kappa) - \frac{1}{2} \mathcal{M}(1/2, 3/2, \kappa) - 1'$$

where $\mathcal{M}$ is the confluent hypergeometric function. The order parameter could be varied to produce geometries between fully coherent (OP = 1) and fully dispersed (OP = 0) orientations. The two remaining environments were designed to represent diffusion in damaged neural tissue ($C_1$) and CSF ($C_{CSF}$). The diffusion in these environments was assumed to be isotropic, with a domain mean diffusivity of $MD_k = 1.7$ and 3.0 µm²/ms in $C_1$ and $C_{CSF}$, respectively.

Damaged WM was simulated by gradually replacing $C_0$ with $C_1$. This was done in four geometries; the first three included one, two and three coherent (OP = 1) and orthogonal $C_2$ components, and the last contained one $C_3$ component with randomly oriented domains (OP = 0). The isotropic component replaced one anisotropic component while the remaining anisotropic components were unaltered. For example, in the case of two crossing fibers ($C_{x1}$ and $C_{x2}$), the damaged anisotropic component $C_{a1}$ had a volume fraction $f_{a1}$. Initially, $f_{a1}$ made up half the volume, but was gradually reduced to zero, and the fraction lost in $C_{a1}$ was replaced by $C_0$, i.e., $f_{a1} = 1/2 \rightarrow 0$, and $f_{a1} = 1/2 \rightarrow f_{a1}$. During this process the fraction of $C_{a1}$ was constant ($f_{a2} = 1/2$).

The response to increasing radial diffusivity, mimicking demyelination, was simulated in a coherent $C_0$ component (OP = 1), where the radial diffusivity was increased from its starting value until it exhibited no anisotropy ($RD_k = 0.2 \rightarrow 1.7$ µm²/ms). Effects of orientation dispersion were investigated using a single $C_0$ component with variable amount of dispersion, from dispersed to coherent (OP = 0 → 1). The effect of the crossing angle between two coherent $C_0$ components was simulated by varying the angle from a parallel to a perpendicular geometry ($\varphi = 0 \rightarrow 90^\circ$). Finally, the effects of CSF contamination were simulated by gradually replacing a coherent $C_0$ component (OP = 1) with $C_{CSF}$ ($f_1 = 1 \rightarrow 0$, and $f_{CSF} = 1 \rightarrow f_{CSF}$). In all cases, the effects of noise were simulated for five equidistant points along each process by adding Rice-distributed noise to the signal (Sijbers and den Dekker, 2004). The signal was generated in accordance with the imaging protocol, i.e., using the same b-values, number of directions and parameterization, at a SNR of 20. The model was regressed onto 1000 realizations of the noisy signal to render a reliable median and interquartile range of the parameters.

Results

In vivo experiments

Maps of FA, µFA and OP for one healthy volunteer are shown in Fig. 4. As expected, the µFA is high in regions comprised of WM and lower in GM. Most notably, the FA and µFA maps differ in regions where a high orientation dispersion is expected, for example, in crossing WM and the interface between WM pathways, in accordance with Lawrenz and Finsterbusch (2014). Another prominent difference can be seen in the GM where FA is close to zero, whereas µFA indicates that the GM contains detectable microscopic anisotropy. Fig. 5 shows the parameter distribution in the CC, CST, CR, THA and CSF, and the powder averaged signal originating from a single voxel in each region. As expected for WM tissue, the departure from monoexponential attenuation was

![Image](https://example.com/image.png)
smaller for the isotropic encoding than the anisotropic encoding. The THA exhibited a relatively high isotropic variance, but the presence of microscopic anisotropy is clearly visible from the separation of the two signal curves. In the CSF, the signal was attenuated below 5% of its initial value, and it is apparent that the fitting would detect a false variance if high b-value data was not excluded. The resulting parameterization of the signal seen in Fig. 5 was: \( \mu_{\text{FA}} = 0.98, 1.03, 0.96, 0.76, \) and 0.00; \( \mu_{\text{MD}} = 0.91, 0.84, 0.89, 1.60, \) and 2.95 \( \mu^2/\text{ms}^2; \) and \( \mu_{\text{V}} = 0.07, 0.00, 0.01, 1.66, \) and 0.01 \( \mu^2/\text{ms}^2, \) and \( \mu_{\text{V}} = 0.57, 0.66, 0.51, \) 0.65, and 0.00 \( \mu^2/\text{ms}^2 \) in the CC, CST, CR, THA and CSF, respectively.

The voxel-wise correlation between \( \mu_{\text{FA}}, \mu_{\text{OP}} \) and FA was presented in Fig. 6. The relation between FA and \( \mu_{\text{FA}} \) resembles the relation between the corresponding parameters reported by Jespersen et al. (2013) in that high FA entails high \( \mu_{\text{FA}}, \) although not vice versa. The correlation between \( \mu_{\text{FA}} \) and FA was found to exhibit two distinct modes, which were separated by introducing an arbitrary threshold at the shoulder of the distribution (\( \mu_{\text{FA}} = 0.8 \)). The interval containing high values of \( \mu_{\text{FA}} \) was found to correspond well to regions of WM (\( \mu_{\text{FA}} > 0.8, \) red outline in Fig. 6) and the low \( \mu_{\text{FA}} \) was found in a mixture of peripheral WM, GM and CSF (\( \mu_{\text{FA}} < 0.8, \) white outline in Fig. 6). In the WM region, a strong correlation was found between OP and FA (\( r^2 = 0.9 \)), while only weak correlations were found between \( \mu_{\text{FA}} \) and OP (\( r^2 = 0.1 \)), and between \( \mu_{\text{FA}} \) and FA (\( r^2 = 0.4 \)). No relevant correlations were found in the peripheral region (all \( r^2 < 0.3 \)).

The investigation of the parameter distribution in the group of healthy volunteers is summarized in Table 1. All parameter mean values, except the MD and \( V_i \), were found to have significantly different mean values in the three WM ROIs. This was expected for the FA since the ROIs include both coherent and crossing WM tissue. The \( \mu_{\text{FA}} \) was also found to differ significantly between the three regions, albeit at a much smaller effect size compared to the FA. The group level variability detected in MD and \( V_i \) indicated that the absence of significance is likely due to a small effect size and large variance, respectively.

The anisotropy parameters measured in the two tumor types are presented in Fig. 7, and corresponding microphotos of the excised tumors are presented in Fig. 8. The meningioma tissue exhibited a low voxel scale anisotropy (mean ± standard deviation, FA = 0.19 ± 0.06) and high microscopic anisotropy (\( \mu_{\text{FA}} = 0.88 ± 0.08 \)). Likewise, the glioblastoma tissue exhibited low voxel scale anisotropy (FA =

<table>
<thead>
<tr>
<th>Region</th>
<th>MD [\mu^2/\text{ms}]</th>
<th>FA</th>
<th>\mu_{\text{FA}}</th>
<th>OP</th>
<th>\mu_{\text{OP}}</th>
<th>V_i [\mu^2/\text{ms}^2]</th>
<th>V_\text{csf} [\mu^2/\text{ms}^2]</th>
</tr>
</thead>
<tbody>
<tr>
<td>THA</td>
<td>1.09 ± 0.20</td>
<td>0.56 ± 0.11</td>
<td>0.56 ± 0.05</td>
<td>1.00 ± 0.06</td>
<td>0.31 ± 0.04</td>
<td>0.64 ± 0.04</td>
<td>0.38 ± 0.04</td>
</tr>
<tr>
<td>CC</td>
<td>0.82 ± 0.09</td>
<td>0.87 ± 0.03</td>
<td>0.97 ± 0.01</td>
<td>0.93 ± 0.01</td>
<td>0.26 ± 0.02</td>
<td>0.64 ± 0.04</td>
<td>0.27 ± 0.03</td>
</tr>
<tr>
<td>CST</td>
<td>0.50 ± 0.12</td>
<td>0.56 ± 0.19</td>
<td>0.65 ± 0.07</td>
<td>0.57 ± 0.07</td>
<td>0.54 ± 0.40</td>
<td>0.30 ± 0.21</td>
<td>0.17 ± 0.05</td>
</tr>
<tr>
<td>CR</td>
<td>0.12 ± 3</td>
<td>0.5 ± 3</td>
<td>32 ± 5</td>
<td>24 ± 5</td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

\( ^\dagger \text{ANOVA shows significant difference between parameter mean values (p < 0.05/6) in the CC, CST and CR.} \)

Table 1: Diffusion parameters (group mean ± standard deviation) in four ROIs in the group of healthy volunteers (n = 8). The ANOVA indicated significantly different mean values in the CC, CST and CR for all parameters except MD and \( V_i \). Note that the number of voxels in each ROI (#Vox) is shown but was not included in any tests.
0.07 ± 0.05). However, it exhibited markedly lower microscopic anisotropy compared to the meningioma ($\mu FA = 0.39 ± 0.22$). Although both tumors exhibited low FA values, the FA in the meningioma was elevated compared to the glioblastoma, indicating that the tissue is organized enough to create a weak but detectable diffusion anisotropy on the voxel scale. The high vs. low microscopic anisotropy in the meningioma and glioblastoma was corroborated by the histological examination of the two tumors, shown in Fig. 8. The histological examination of the meningioma demonstrated a dense fascicular pattern of growth with elongated tumor cells, consistent with low FA and high $\mu FA$; and a more loose assemblage of rounded cells of variable size along with patchy areas of necrosis in the glioma, consistent with both low FA and low $\mu FA$.

**Simulation experiment**

Figs. 9 and 10 showcase how the FA and $\mu FA$ are altered when the underlying diffusion profiles are manipulated. When a coherent anisotropic component was replaced by an isotropic component (Fig. 9A), the FA decreased approximately linearly as a function of the isotropic tissue fraction. In the same system, the $\mu FA$ followed...
a similar pattern, but had a less pronounced initial slope indicating that the μFA is overestimated when the distribution of diffusion coefficients contains both isotropic and anisotropic components. In the absence of noise, both parameters approached zero for purely isotropic systems. In the crossing geometry, where one anisotropic component was replaced by an isotropic component (Fig. 9B), the FA first decreased due to the relatively rapid increase of the isotropic component. However, when a majority of the receding component had been removed \((f_i > 1/2)\), the FA instead increased due to the dominance of the remaining anisotropic component. By contrast, μFA decreased strictly. This demonstrates a case where μFA may exhibit superior sensitivity and specificity over FA, since the direction of the effect is constant. Further, the effect size is larger for μFA since it is not confounded by the same counteracting mechanisms. Similar results are shown for a triple crossing geometry (Fig. 9C). In this case the FA started at a low value because the tissue was macroscopically isotropic with its three orthogonal fiber populations, and increased as one of the fiber populations was replaced by isotropic tissue. Again, the positive direction of the effect, caused by the reduction in orientation dispersion, may be confounding. By contrast, μFA reflected only the presence of microscopic anisotropy and responded as expected to the simulated damage. In the case of damage in randomly oriented microdomains (Fig. 9D), the macroscopic anisotropy is zero, rendering FA insensitive to any changes in tissue microstructure while the μFA reflects the amount of microscopic anisotropy that is lost.

The effect of gradually increasing domain radial diffusivity, resulted in similar effects for FA and μFA (Fig. 10A). However, as the system approaches isotropic conditions, the uncertainty in the μFA increases considerably. Fig. 10B shows how dispersion influences the FA, while the μFA is constant. A similar pattern is seen when simulating crossing fibers with varying crossing angles (Fig. 10C). As expected the FA was highest when the two fiber structures were parallel and had its lowest value when they were perpendicular. These results show the potential benefits of quantifying a measure for anisotropy that is not sensitive to confounds such as crossing, bending, fanning, and kissing fiber geometries. Finally, the effects of CSF contamination exhibit similar effects as the simulated damage in a single coherent WM system (compare Figs. 9A and 10D). This simulation highlights the overestimation of μFA due to multiple isotropic components. Generally, the μFA is increasingly susceptible to noise as the simulated systems approach zero microscopic anisotropy, resulting in reduced accuracy.
Discussion

In this study we present the first implementation of qMAS for the purpose of probing the microscopic anisotropy in vivo on a clinical MRI system. The parameters $\mu^2$FA and OP, as well as conventional DTI parameters FA and MD, were quantified in healthy subjects and in two different types of tumor tissue. Unlike the voxel scale anisotropy, measured in terms of the FA, the microscopic anisotropy measured by $\mu^2$FA was relatively homogeneous in large portions of the WM. This finding is in agreement with other studies that have aimed to remove effects of orientation dispersion from the quantification of local anisotropy (Jensen et al., 2014; Jespersen et al., 2013; Lawrenz and Finsterbusch, 2013, 2014). The notion that FA is sensitive to local orientation dispersion is supported by the strong correlation found between the FA and OP (Fig. 6). However, the three WM regions chosen for analysis exhibited small but statistically significant differences also in $\mu^2$FA (Table 1), indicating that orientation dispersion is not the only difference between these regions. This could possibly be explained by varying levels of domain anisotropy, for example, caused by variable axonal packing density.

In the tumor tissue, FA was generally low, which indicated that the meningioma and the glioblastoma were approximately isotropic on the voxel scale. By contrast, the $\mu^2$FA was able to reliably differentiate between the two tumors, and indicated that microscopic diffusion anisotropy was more pronounced in the meningioma than the glioblastoma. Thus, the information provided by both FA and $\mu^2$FA was instrumental in predicting the tumor cell structures which were later confirmed by the histological exam (Fig. 8).

To elucidate some of the underlying mechanisms that affect FA and $\mu^2$FA, simulations of different micro-environments visualized the parameters as a function of several relevant processes. For example, in the case of increased radial diffusivity of parallel fibers, the responses in FA and $\mu^2$FA are approximately equal, meaning that the two representations of anisotropy share a common interpretation. On the other hand, scenarios that include any form of orientation dispersion demonstrate prominent differences between FA and $\mu^2$FA. For example, the combination of two and three orthogonal anisotropic components (Figs. 9B and C) were used to reproduce the effects of selective atrophy in a crossing WM geometry, as reported by Douaud et al. (2011), where the effect direction in FA was found to be positive in a damaged region of crossing WM. The simulations also illuminated the bias that arises when $\mu^2$FA is quantified in systems that violate the assumptions used in the parameterization, e.g., in complex mixtures of anisotropic and isotropic tissue. Although these scenarios invalidate the $\mu^2$FA as a direct metric of the microscopic anisotropy, it is worth noting that it retains sensitivity to the relevant effect and does so in a more consistent manner than the FA.

Although the comparison between FA and $\mu^2$FA showcases the effects of orientation dispersion as a confounder for FA, it does not invalidate previous studies that employ FA as a biomarker. Instead, the origin of the effect can be better understood, possibly allowing an improved interpretation of the FA and its relation to the microstructural integrity. We expect that $\mu^2$FA may not only contribute to the investigation of complex WM geometries, but also in detecting microscopic anisotropy in tissues that are approximately isotropic on the voxel scale, for example, in GM (McNab et al., 2013; Truong et al., 2014). Further, the $\mu^2$FA and OP may provide complementary information to the FA and tensor shape analysis previously used in the differentiation of classic and atypical meningioma (Toh et al., 2008), detection of fibroblastic meningioma (Tropine et al., 2007), and in the preoperative estimation of tumor consistency (Kashimiru et al., 2007), by removing the confounding effects of orientation dispersion which are otherwise ignored.

It is important to stress that the signal acquired with conventional anisotropic encoding used in this study is identical to that needed for DKI analysis. However, because DKI makes no effort to distinguish between the origins of the diffusion kurtosis (herein referred to as variance in diffusion coefficients) it is not directly associated to microscopic anisotropy. The framework presented here is also related to the dPFG-methods employed by Jespersen et al. (2013) and Lawrenz and Finsterbusch (2014). In terms of the analysis presented here, dPFG encoding can be described with an encoding tensor which renders a signal that is sensitive to a weighted sum of $V_i$ and $V_o$, where the weighting depends on the direction of the encoding blocks (Westin et al., 2014). It appears that the framework based on qMAS combined with anisotropic encoding probes the $\mu^2$FA more directly and may therefore provide a faster technique for measuring microscopic anisotropy compared to the dPFG methods. Finally, we note that the implementation and use of qMAS is no more complicated than a similar DKI protocol. Other techniques that take orientation dispersion into account include, for example, NODDI which quantifies the magnitude of fiber dispersion and the neurite density (Zhang et al., 2012). From this information it is possible to calculate a parameter analogous to the $\mu^2$FA. However, like DTI and DKI, the NODDI technique cannot distinguish between randomly oriented anisotropic domains and multiple isotropic components. Another drawback of model-based approaches, such as NODDI, is the demand for a priori assumptions about the tissue that is investigated, which may limit their use in abnormal tissues such as tumors.

In the present study, several factors affected the accuracy, i.e., the trueness and precision, of the estimated $\mu^2$FA. The imaging protocol features a long echo time which impacted the SNR and thus also the precision of $\mu^2$FA. Sufficient SNR for a robust signal parameterization was achieved by increasing the voxel size. Consequently, this increased the amount of PVE, especially in tissue interfacing with CSF, thereby reducing the trueness in such regions. Note that the present protocol was designed to test the validity of the suggested model by acquiring a densely sampled signal. However, the experimental design can be adjusted to allow whole brain coverage at feasible acquisition times by optimizing the acquisition protocol (Alexander, 2008). Further, a relatively low number of encoding directions were acquired, which may have reduced the trueness by introducing a weak directional dependency in the powder averaged signal, although simulations (data not shown) indicate a negligible $\mu^2$FA bias even for highly anisotropic tissue. A further limitation of $\mu^2$FA is that it may suffer from low accuracy when the model assumptions are violated or when investigating tissue with little or no microscopic anisotropy. The effects of such unfavorable conditions are demonstrated in the simulations (Figs. 9 and 10). The reduced accuracy in tissue with low anisotropy ($\mu^2$FA < 0.4) can be understood by considering Eq. (7) for $V_o$ approaching zero; where the restriction on $V_o$ to be positive may reduce trueness, and low levels of variance in $V_o$ will render a poor precision in $\mu^2$FA. Thus, it is likely that the $\mu^2$FA calculated in the glioblastoma exhibited a positive bias since the histological exam of the glioblastoma found few anisotropic structures (Fig. 8). Although the accuracy of the estimated $\mu^2$FA in the glioblastoma may be poor, the tumors could be reliably differentiated based on the difference in their microscopic anisotropy. Finally, a limitation may be that the assumption of Gaussian diffusion is not valid, i.e., that the signal attenuation may be dependent on diffusion time. We do not expect this to be the case in white matter for the current diffusion time regime (Nilsson et al., 2009, 2013). However, tumor tissue may contain larger cell structures, which could make $\mu^2$FA dependent on experimental parameters. This is a topic that deserves further attention, especially since qMAS exhibits an anisotropic time dependency due to the varying speed of the q-vector through q-space (Fig. 3).

Conclusion

This study demonstrates the feasibility of mapping the microscopic anisotropy of the brain in vivo in terms of the $\mu^2$FA. The Results suggest that the contrast found in conventional FA maps is strongly modulated by the orientation dispersion of the anisotropic domains contained within each imaging voxel. By contrast, our analysis quantifies the microscopic anisotropy and orientation dispersion separately in terms of the $\mu^2$FA and OP. Unlike the conventional FA derived from DTI, $\mu^2$FA may
therefore provide a robust biomarker that probes the relevant diffusion anisotropy even in complex WM configurations. The potential benefit of μFA was demonstrated in two brain tumors. Although both tumors appeared isotropic on the voxel scale, the μFA could be used to distinguish between them based on their microscopic anisotropy. Additionally, simulations of complex tissue microstructures suggested that μFA exhibits a more intuitive interpretation than FA.

We predict that the combination of FA, μFA and OP can be useful in clinical and research applications, by enabling detection of microstructural degeneration in complex neural tissue, detection of fibrous tissue in tumors for pre-surgical classification of consistency, and quantification of microscopic anisotropy in macroscopically isotropic tissue.

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Conflict of interest statement

SL, DT and MN declare patent applications in Sweden (1250453-6 and 1250452-8), USA (61/642 594 and 61/642 589), and PCT applications of interest statement

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