Neighbourhood social participation and women's use of anxiolytic-hypnotic drugs: a multilevel analysis.

Johnell, K; Merlo, Juan; Lynch, J; Blennow, G

Published in:
Journal of Epidemiology and Community Health

DOI:
10.1136/jech.58.1.59

2004

Citation for published version (APA):
Neighbourhood social participation and women’s use of anxiolytic-hypnotic drugs: a multilevel analysis

K Johnell, J Merlo, J Lynch, G Blennow

Study objectives: To identify and quantify a hypothesised collective effect of the neighbourhood on individual use of anxiolytic-hypnotic drugs (AHD). To analyse the general impact of neighbourhood social participation on use of AHD, adjusting for individual characteristics.

Design: Cross sectional analysis performed by multilevel logistic regression with women at the first level and neighbourhoods at the second level.

Setting: Malmö (250,000 inhabitants), Sweden.

Participants: 15,456 women aged 45 to 73, residing in 95 neighbourhoods in Malmö, who took part in the Malmö diet and cancer study (1991–1996).

Main results: The prevalence of AHD use was 5.5% in the study sample. Overall, 1.7% of the total individual differences in the propensity for using AHD were explained by the neighbourhood level. This percentage, however, differed between different individuals. Low level of social participation in the neighbourhood was associated with higher probability of AHD use (OR = 3.10 [95% CI 1.51 to 6.41]), independently of individual age, low social participation, low educational level, and living alone. This association was reduced (OR = 2.01 [95% CI 0.97 to 4.14]) after the additional accounting for individual disability pension, low self rated health, stress, and medication for somatic disorders.

Conclusions: The neighbourhood level of social participation seems to affect individual use of AHD, possibly through individual characteristics. However, neighbourhood boundaries play a minor part in understanding individual AHD use in the city of Malmö.

METHODS

The Malmö diet and cancer study

All data used in this study originate from The Malmö diet and cancer study (MDCS), a prospective cohort study performed in the city of Malmö in southern Sweden. The city had a population of about 250,000 inhabitants in 1995 and was administratively divided into 110 neighbourhoods. In this study, we included 95 neighbourhoods, leaving out 15 neighbourhoods that each had less than 20 participants in the MDCS. We set an arbitrary limit of 20 participants to reduce bias regarding the measurement of information about the neighbourhoods.

People were requested to participate in the MDCS by letters of invitation, information through advertisements in the local media and collaboration with large employers in Malmö. In total, letters of invitation provided 80% of the participants.

The 17,388 women, aged 45 to 73 years, who participated in the MDCS cohort, represented 41% of all women born 1923–1950 living in Malmö during the baseline period 1991–1996. Of these women, 89% (15,456 of 17,388) were included in this study. Reasons for exclusion were lack of information on drug use, incomplete information on other variables studied, and living in neighbourhoods with less than 20 participants in the MDCS.

Baseline examination occurred during 1991–1996. A self administered questionnaire and a seven day personal diary were used to obtain information on relevant characteristics of the women, including use of medication. The participants

Abbreviations: AHD, anxiolytic-hypnotic drugs; MDCS, Malmö diet and cancer study; ICC, intraclass correlation; VPC, variance partition coefficient; RIGLS, restricted iterative generalised least squares; PQL, penalised quasilikelihood
completed both information sources at home within one to two weeks between the first and the second consecutive baseline visit to the project office. All participants in this study gave information in both the questionnaire and the diary. Information on how many women had the questionnaire and the diary administered during 1991–1996 is presented in a previous paper.11

The ethical committee at the Medical Faculty of Lund University approved the study proposal of the MDCS and all participants gave signed informed consent. A detailed description of the design and aims of the cohort study is given elsewhere.11

Outcome
We obtained information on medication use from both the self-administered questionnaire and the seven day personal diary. In the self-administered questionnaire, women listed brand name(s) of the drug(s) they had used during the past two weeks. In the seven day personal diary they daily indicated the medicines they used during the week. Thereafter, all brand names were classified in accordance with the 1997 version of the Anatomical Therapeutic Chemical classification system12–14 (ATC-97). Use of anxiolytic-hypnotic drugs (AHD) was defined by the codes N05B (anxiolytics), N05C (hypnotics and sedatives) and N03AE (benzodiazepine derivates).

Explanatory variables
Age was aggregated into four groups: 45–49 (reference), 50–59, 60–69, and 70–73 years. Individual social participation was defined by possible involvement in 13 formal or informal activities (study circle/course at work place, other study circle/course, union meeting, meeting of other organisations, theatre/cinema, arts exhibition, church, sports event, letter to editor of a newspaper/journal, demonstration, night club/entertainment, large gathering of relatives, private party), which the respondent might have participated in during the previous 12 months.15 Items were summed, and participants with involvement in three or fewer activities (lowest quartile) were classified as having low social participation.

Low neighbourhood social participation was assessed by the proportion of persons in the neighbourhood who were classified as having individual low social participation. Less than nine years of formal education was defined as low educational level. Participants also reported whether they were living alone and whether they felt psychological stress outside the work place. All women aged less than 65 years reporting retirement were considered to have a disability pension. An ordinal scale ranging from 1 (“worst possible”) to 7 (“best possible”) determined self rated health (SRH). Low self rated health was defined as a value of ≤4 on this scale. Medication for somatic disorders was defined as use of any pharmaceutical drug except those in the ATC-group N (nervous system).

Statistical analysis
We analysed dichotomous AHD use by multilevel logistic regression analysis16–18 with women at the first level and neighbourhoods at the second level. We constructed four models. The first model i, called the empty model, was without explanatory variables (that is, simple component of variance analysis). The second model ii contained age together with one other variable at a time. The third model iii was created to observe whether social participation at the neighbourhood level affected individual AHD use, over and above the individual variables age, low educational level, low social participation, and living alone. The fourth model iv included all the studied variables.

Fixed effects
We observed the association (the slope of the regression) between individual use of AHD and different individual variables, and the neighbourhood variable percentage of people with low social participation. The β coefficients (standard errors) were converted into odds ratios (95% confidence intervals).

Random effects
We examined whether the neighbourhood environment affected the general individual propensity of using AHD, over and above individual characteristics. This possible contextual effect was measured by the intraclass correlation (ICC), or in this case, the intra-neighbourhood correlation. The ICC represented the percentage of the total variance in the propensity of using AHD that was related to the neighbourhood level and was also used as a measure of clustering of AHD use in the neighbourhoods. The ICC is a simple form of the “variance partition coefficient” (VPC).19 For dichotomous variables, for example, use of AHD, ICC can be calculated according to the formula used by Snijders20 (pages 223–6):

\[
ICC = \frac{V_n}{V_n + \sigma^2}
\]

\[V_n = \text{neighbourhood variance}\]

Calculated in this way, the ICC represents the average ICC for all individuals.19 20

A high ICC in the empty model would indicate high clustering of medication use in the neighbourhoods and strong neighbourhood effect on individual medication use. A low ICC, on the other hand, would express the existence of a weak neighbourhood influence on individual medication use. In separate age adjusted models, we performed several multilevel logistic regression analyses with the individual variables (that is, low social participation, low educational level, living alone, disability pension, low self rated health, self reported stress, missing information on stress, and medication for somatic disorders) as dependent variables.

Analyses of the ICC permitted quantification of the degree to which these individual socioeconomic and health related factors were clustered within the neighbourhoods. In model ii we analysed cross level interactions by letting the slopes of the association between AHD use and the individual variables be random at the neighbourhood level. Random slope variance would indicate that the individual level association was modified by the neighbourhood level (that is, cross level interaction).

Parameters were estimated using the restricted iterative generalised least squares (RIGLS) and penalised quasilikelihood (PQL). Extra-binomial variation was explored systematically in all models and we found no evidence for under-dispersion or over-dispersion. The MLwiN version 1.1 software package21 was used to perform the analyses.

RESULTS
The prevalence of AHD use was 5.5% in the study sample. Of all the 18 different AHD used by the participants, flunitrazepam was the most commonly used drug (1.3%), followed by oxazepam (1.3%), diazepam (0.9%), nitrazepam (0.8%), and propiomazine (0.5%). AHD users had an impaired health and socioeconomic profile, with lower social participation (table 1). The median (1st–3rd quartile) neighbourhood low social participation was 29.3% (22.9%–36.7%). Table 1 shows that as the percentage of individuals with low social participation within the neighbourhoods increased from...
low (group 1) to high (group 4), the women’s health, behaviour and socioeconomic profiles were impaired. Table 1 indicates that, in average, there was a modest neighbourhood clustering of AHD use (ICC = 1.7%). Table 1 also suggests the existence of socioeconomic segregation in the neighbourhoods as illustrated by the ICC for the individual variables living alone (ICC = 20.3%), low educational level (ICC = 14.5%), low social participation (ICC = 7.0%), and disability pension (ICC = 8.9%). In contrast, the clustering was very small (ICC around 1% or less) in relation to health related variables.

**Fixed effects**

Neighbourhood low social participation was associated with higher individual probability of use of AHD (table 2). Model iii shows that, independently of each other, both individual and neighbourhood social participation were associated with AHD use. After inclusion of all the individual variables in

### Table 1

| Characteristics of the 15 456 women residing in 95 of the 110 neighbourhoods in the city of Malmö. Values are percentages unless otherwise indicated |
|---------------------------------|---------------|
| **Use of AHD** | **Neighbourhood quartile groups of increasing rate of low social participation (mean %)** |
|               | No | Yes | Group 1 | Group 2 | Group 3 | Group 4 |
| Use of AHD    | 1.7 | 94.5 | 5.5 | 4.4 | 5.3 | 5.6 | 6.9 |
| Age           | 6.0 | 57.3 | 60.3 | 56.2 | 57.0 | 57.3 | 58.1 |
| Low social participation | 7.0 | 29.5 | 42.6 | 19.1 | 25.3 | 36.2 | 45.5 |
| Low educational level | 14.5 | 28.7 | 42.6 | 18.8 | 23.4 | 46.0 | 61.2 |
| Living alone  | 20.3 | 27.7 | 41.7 | 17.0 | 21.5 | 29.4 | 38.1 |
| Disability pension | 8.9 | 4.9 | 12.9 | 3.2 | 4.2 | 5.3 | 9.0 |
| Low self rated health | 1.3 | 27.8 | 61.4 | 24.1 | 28.1 | 31.4 | 35.2 |
| Self reported stress | 0.8 | 30.9 | 54.1 | 32.4 | 31.1 | 28.6 | 31.8 |
| Not answering the question on stress | 0.1 | 5.2 | 10.2 | 4.6 | 5.7 | 4.8 | 7.1 |
| Medication for somatic disorders | 0.2 | 61.7 | 83.0 | 63.2 | 64.4 | 62.6 | 61.3 |

*Intraclass (that is, intra-neighbourhood) correlation (ICC). The ICC measures the percentage of the total variance of the individual variable that is related to the neighbourhood level. It gives information about the clustering of the individual characteristics in the neighbourhoods.*

### Table 2

<table>
<thead>
<tr>
<th>ICC (%)</th>
<th>Neighbourhood intercept variance (standard error)</th>
<th>Neighbourhood slope variance (standard error)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Model i (empty model)</td>
<td>1.7</td>
<td>0.056 (0.025)</td>
</tr>
<tr>
<td>Age adjusted models</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Age</td>
<td>Ref</td>
<td></td>
</tr>
<tr>
<td>45–49</td>
<td>1.94</td>
<td>(1.14 to 2.45)</td>
</tr>
<tr>
<td>50–59</td>
<td>2.95</td>
<td>(2.34 to 3.72)</td>
</tr>
<tr>
<td>60–69</td>
<td>2.98</td>
<td>(2.27 to 3.92)</td>
</tr>
<tr>
<td>70–73</td>
<td>1.54</td>
<td>(1.33 to 1.78)</td>
</tr>
<tr>
<td>Low social participation (yes v no)</td>
<td>0.93</td>
<td>(0.79 to 1.10)</td>
</tr>
<tr>
<td>Low educational level (yes v no)</td>
<td>1.70</td>
<td>(1.45 to 1.99)</td>
</tr>
<tr>
<td>Living alone (yes v no)</td>
<td>3.02</td>
<td>(2.42 to 3.77)</td>
</tr>
<tr>
<td>Disability pension (yes v no)</td>
<td>4.29</td>
<td>(3.67 to 5.01)</td>
</tr>
<tr>
<td>Stress</td>
<td>Ref</td>
<td></td>
</tr>
<tr>
<td>Self reported stress (yes v no)</td>
<td>3.19</td>
<td>(2.69 to 3.79)</td>
</tr>
<tr>
<td>Missing information on self reported stress (yes v no)</td>
<td>2.95</td>
<td>(2.16 to 4.04)</td>
</tr>
<tr>
<td>Medication for somatic disorders</td>
<td>2.91</td>
<td>(2.38 to 3.56)</td>
</tr>
<tr>
<td>(yes v no)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Neighbourhood rate of low social participation</td>
<td>5.02</td>
<td>(2.39 to 10.55)</td>
</tr>
<tr>
<td>Model iii</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Low social participation (yes v no)</td>
<td>1.56</td>
<td>(1.34 to 1.81)</td>
</tr>
<tr>
<td>Low educational level (yes v no)</td>
<td>0.81</td>
<td>(0.70 to 0.95)</td>
</tr>
<tr>
<td>Living alone (yes v no)</td>
<td>3.10</td>
<td>(1.35 to 1.86)</td>
</tr>
<tr>
<td>Neighbourhood rate of low social participation</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Model iv</td>
<td></td>
<td></td>
</tr>
<tr>
<td>All variables in the model</td>
<td>2.01</td>
<td>(0.97 to 4.14)</td>
</tr>
</tbody>
</table>
model iv, the association between the neighbourhood variable and AHD use was weakened and slightly uncertain using 95% confidence intervals.

**Random effects**

Model i (empty model) revealed that only a relatively small percentage of the total individual variability in the propensity of using AHD was attributable to the neighbourhood level (ICC = 1.7%).

The random slopes variance analysis suggested that the neighbourhood modified the association between certain individual level variables and AHD use (figs 1 and 2). Figure 2 shows the predicted probability of using AHD in function of self reported stress. Both the women who reported stress (ICC = 2.1%) and those who did not answer this question (ICC = 11.1) had a higher probability of AHD use. This association was, however, different for different neighbourhoods.

**DISCUSSION**

**Main findings**

Our results suggest that the level of low social participation in the neighbourhood may determine AHD use among women, over and above individual social participation, educational level, and living alone. This association was, however, reduced after additional adjusting for self reported stress, self rated health, disability pension, and use of medication for somatic disorders.

The observed weakening of the association between neighbourhood social participation and women’s use of AHD after we adjusted for the individual health variables may be an expression for compositional confounding—that is, women with impaired health move to certain neighbourhoods. However, the observed reduction of the association may instead be telling us that these individual level variables were in the pathway between neighbourhood social environment and use of AHD. For instance, living in a deprived neighbourhood with low social participation may negatively affect individual social participation and health related behaviour, which in turn could have an impact on medication use. The neighbourhood environment may influence AHD use through diverse mechanisms like accessibility to health services and local practice habits of practitioner, normative attitudes towards medication use, transmission of health information, health behaviour norms, and social control of deviant health related behaviour. Our results suggest that general practitioners could be aware of the association between neighbourhood low social participation and AHD use. At a political level, it may be emphasised

**Policy implications**

- Our results suggest that general practitioners could be aware of the association between neighbourhood low social participation and AHD use.
- At a political level, it may be emphasised that low social participation in neighbourhoods seems to go hand in hand with higher use of AHD.
that low social participation in neighbourhoods seems to go hand in hand with higher use of AHD.

Although neighbourhood social participation was associated with AHD use, individual AHD use was on average rather homogeneous all over the city of Malmö. That is, only 1.7% of the individual differences in the probability of using AHD seemed to have contextual reasons. In other words, there existed a collective phenomenon influencing AHD use, but it explained a minor part (only 1.7%) of the individual variation in the propensity for using AHD.

In contrast with medication use, there existed a stronger clustering for some socioeconomic variables, like for example living alone, low educational achievement, and disability pension. The results of cross sectional multilevel analyses often demonstrate a higher neighbourhood clustering of socioeconomic and behavioural outcomes than of biological and medical variables. Probably, socioeconomic and behavioural variables are more susceptible to current contextual influences than life course determined biomedical variables. If contextual factors do affect biomedical outcomes, these effects need to be investigated by longitudinal, rather than cross sectional, multilevel analysis.

The random slope variance analysis suggested that the geographical units modified the association between certain individual level variables and AHD use. This impact of the neighbourhood on AHD use, as expressed by the ICC, was larger for stressed women and for those who abstained from answering this specific question on stress, than for women without stress. That is, in average, the neighbourhood level explained a modest part of the individual differences in the probability of using AHD, but this average measure hid larger neighbourhood effects for certain groups of women.

At this point, we will comment on the fact that the women with missing information on stress reported very similar AHD use as women reporting stress (that is, 9% and 10% respectively). It is possible, therefore, that women with missing information on this question were in fact women suffering from stress.

**Limitations: selection bias, information bias, and confounding**

The selection of geographical units should not be a source of selection bias, as we only left out the 15 least populated neighbourhoods. The participation rate (median; first–third quartile) in the MDCS in the different neighbourhoods was low (42%; 32%–50%). Hence, the cohort may not be representative for the whole female population, which may reduce the external validity of our results. However, participants could be regarded as fairly representative of the general population, at least in relation to the main socio-demographic variables.

Information on drug use was self reported, which seems to be a valid method for measuring current drug use. The reliability (test-retest stability) of the social participation variable, as assessed in a previous paper, was rather high (x = 0.70). Unfortunately, we had no information about how long the women had lived in their respective neighbourhoods, which would have given information about the exposure time of the specific neighbourhood rate of low social participation.

**Conclusions**

Our study gives empirical support to the idea that contextual factors related to aspects of the neighbourhood environment contribute to differences in individual behaviour, such as medication use. In concrete, our results suggest that a low level of social participation in the neighbourhood, a main aspect within the concept of social capital, seems to be related to individual AHD use, over and above individual social participation, living alone, and educational level. In a multilevel way of thinking, neighbourhood social participation can be placed upstream in the causal pathway determining individual health and it may affect downstream individual characteristics, which in turn influence the use of AHD.

However, neighbourhoods in Malmö are on average rather homogeneous regarding individual use of AHD (that is, the neighbourhood level explained 1.7% of the differences in the probability of using AHD). Nevertheless, neighbourhoods play a larger part in explaining AHD use for certain groups of women, like those reporting stress. This heterogeneity needs to be considered when analysing contextual effects on individual health.

**ACKNOWLEDGEMENTS**

We wish to express our gratitude to Professor Göran Berglund for his support regarding the Malmö diet and cancer study database and to the NEPI foundation for its support regarding studies in social pharmacoepidemiology. We also wish to thank Min Yang at the Centre for Multi-level Modelling, Institute of Education, London for useful comments about multilevel modelling. This study was financed by the County of Skåne’s Pharmacological council, by an ALF-government grant, Dnr M: E 39 390/98 (Juan Merlo), the NEPI foundation, the Swedish council for working life and social research, the Swedish Medical Research Council, the Swedish Cancer Society and National Institutes of Health (RO1 HL71084-01).

**Authors’ affiliations**

K Johnell, Family Medicine Stockholm, Karolinska Institutet, Huddinge, Sweden

J Merlo, Department of Community Medicine, Malmö University Hospital, Lund University, Malmö, Sweden

J Lynch, Department of Epidemiology, School of Public Health Center for Social Epidemiology and Population Health, and Institute for Social Research, University of Michigan, Ann Arbor, USA

G Blennow, The National Social Insurance Board, Stockholm, Sweden

**REFERENCES**


www.jech.com